Gaussian Process Regression and Emulation STAT8810, Fall 2017

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More on GP Regression

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- A variogram written as a function of ||h|| is said to be *isotropic*.
 Otherwise, it is *anisotropic*.

• The semi-variogram for the powered exponential model is

$$\gamma_{Y}(h;\theta) = c_{Y}(0;\theta) - c_{Y}(h;\theta)$$

where $c_{Y}(h;\theta) = \sigma^{2} exp\left(-\frac{||h||^{\theta_{2}}}{\theta_{1}}\right)$.

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• The empirical semi-variogram is calculated as

$$\hat{\gamma}_{Y}(h) = \frac{1}{2} \operatorname{ave} \left\{ (\overset{\mathsf{Y}}{\boldsymbol{\mathscr{I}}}(x_{i}) - \overset{\mathsf{Y}}{\boldsymbol{\mathscr{I}}}(x_{j}))^{2} : ||x_{i} - x_{j}|| \in T(h), i, j = 1, \dots, n \right\}$$

where T(h) is some tolerance region around h (e.g. $h \pm \Delta$, for Δ small).

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- How do we know if a correlation model is "good" for our data?
- 1. Eyeball test: compare the theoretical semivariogram (at various θ 's or maybe $\hat{\theta}_{MLE}$) versus the empirical semivariogram.
- **2.** Simple test of H_0 : $Z(\cdot)$ has no spatial dependence.[†]

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2. Let $F = \frac{\hat{\gamma}_Y(h_1)}{\hat{\sigma}^2}$ where $\hat{\sigma}^2 = \frac{\sum (Z(x_i) - \hat{\mu})^2}{n-1}$, $\hat{\mu} = \frac{1}{n} \sum Z(x_i)$ and h_1 is the smallest lag from all possible lags h_1, \ldots, h_L .

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- Reject H₀ for |F 1| large. How? Permutation-based test is one distribution-free approach:
- i: Permute the data locations x_{σ(1)},..., x_{σ(n)}
- ii: Recompute F
- iii: Repeat steps i,ii many times
- iv: If the observed F is above the 97.5th percentile or below the 2.5th percentile of the permutation distribution, reject H_0 at the 5% level.

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 - Repeat i,ii many times.
 - Construct variogram "envelopes" from the draws[†] and see if the observed empirical variogram of the dataset lie within the envelope.

```
source("dace.sim.r")
 Generate a 2-D realization
set.seed(99) #remove for a "real" realization
library(rgl) #for nice plots
n=10
x1=x2=seq(0,1,length=n)
design=as.matrix(expand.grid(x1,x2))
l1=list(m1=abs(outer(design[,1],design[,1],"-")))
12=list(m2=abs(outer(design[,2],design[,2],"-")))
1.dez=list(11=11,12=12)
rho=c(0.5, 0.5)
alpha=2
s2=1
se2=0
z=sim.field(l.dez,rho,s2,se2=se2,alpha=alpha,
```



library(geoR)

-----## Analysis of Geostatistical Data ## For an Introduction to geoR go to http://www.leg.ufpr.k ## geoR version 1.7-5.2 (built on 2016-05-02) is now loade ## ------

gsim=cbind(design,z) gsim=as.geodata(gsim) vgram=variog(gsim)

variog: computing omnidirectional variogram

eyefit(vgram) # Explore different models, etc.

0

0

0



Figure 2:



Construct envelope using permutation test.
vgram.env=variog.mc.env(gsim, obj.var = vgram)

variog.env: generating 99 simulations by permutating dat
variog.env: computing the empirical variogram for the 99
variog.env: computing the envelops

plot(vgram, envelope = vgram.env)



distance

##	kappa no	ot used for the gaussian correlation function
##		
##	likfit:	likelihood maximisation using the function optim
##	likfit:	Use control() to pass additional
##		arguments for the maximisation function.
##		For further details see documentation for optim
##	likfit:	It is highly advisable to run this function seven
##		times with different initial values for the par-
##	likfit:	WARNING: This step can be time demanding!
##		
##	likfit:	end of numerical maximisation.

plot(vgram, envelope = vgram.env)



distance

```
# Instead of plotting max/min envelope, plot the 5%
# and 95% quantiles from the simulations
gtmp=gsim
nbins=length(vgram$u)
vgsims=matrix(0,nrow=100,ncol=nbins)
for(i in 1:100) {
    gtmp$data=vgram.env$simulated.data[,i]
    vg=variog(gtmp)
    vgsims[i,]=vg$v
}
```

variog: computing omnidirectional variogram

plot(vgram, envelope = vgram.env)
for(i in 1:nbins) points(vg\$u[i],quantile(vgsims[,i],0.05)
for(i in 1:nbins) points(vg\$u[i],quantile(vgsims[,i],0.95)



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- path-wise properties depend on behaviour of $Z(x) = Z(x, \omega)$ when ω is fixed.
- process-wise properties given as usual specification of mean and covariance.
Definition: Suppose Z(x) is a process with finite second moments. Then, $Z(\cdot)$ is *mean-square continuous* at x_0 if $E[|Z(x)|^2] < +\infty$ and

$$\lim_{x\to x_0} E_{\Omega}\left[|Z(x)-Z(x_0)|)^2\right]=0.$$

Definition: Z(x) has almost surely continuous sample paths on χ provided

$$P(\omega: \forall x_0 \in \chi, z(x) \rightarrow z(x_0) \text{ as } x \rightarrow x_0) = 1.$$

 this one is more abstract. It says that the only events that occur with probability 1 are those where the sample path of the process is continuous.

 Let c_Y(·) be the covariance function of a stationary process Y(·). Then,

$$E\left[(Y(x) - Y(x_0))^2\right] = E\left[Y^2(x) - 2Y(x)Y(x_0) + Y^2(x_0)\right]$$

= 2c_Y(0) - 2c_Y(x - x_0)
= 2(c_Y(0) - c_Y(x - x_0))

So,

$$\lim_{x \to x_0} E\left[\left(Y(x) - Y(x_0)\right)^2\right] = \lim_{x \to x_0} c_Y(0) - c_Y(x - x_0) = 0$$

$$\Leftrightarrow \lim_{x \to x_0} c_Y(x - x_0) = c_Y(0)$$

i.e. the process is mean-square continuous $\forall x_0 \in \chi$ provided $c_Y(\cdot)$ is continuous at the origin. Or similarly, $R(h) \rightarrow 1$ as $h \rightarrow 0$ and R(h) is continuous at origin.

• For GP's, mean-square continuity can be expressed as $E\left[|Z(x_i) - Z(x_j)|^2\right] \to 0 \text{ as } ||x_j - x_i|| \to 0.\dagger$

† Theorem 3.4.1 in Adler: The Geometry of Random Fields (1981).

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- For GP's, mean-square continuity can be expressed as $E\left[|Z(x_i) - Z(x_j)|^2\right] \rightarrow 0$ as $||x_j - x_i|| \rightarrow 0.\dagger$
- It has also been shown that the sample paths of a GP are almost surely continuous if R(h) converges to 1 sufficiently slow:

$$1-R(h) \leq rac{c}{|log(||h||_2)|^{1+\epsilon}} orall ||h||_2 < \delta$$

for some $0 < c < \infty$, some $\epsilon > 0$ and some $\delta < 1$ then $Z(\cdot)$ has a.s. continuous sample paths.

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- It has also been shown that the sample paths of a GP are almost surely continuous if R(h) converges to 1 sufficiently slow:
 - \ddagger If $Z(\cdot)$ is a stationary GRF with correlation function R satisfying

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• In other words, $(1 - R(h))|log(||h||_2)|^{1+\epsilon} \leq c$ is bounded.

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```
c=1
eps=4
h=seq(0,1,length=100)
bound=c/( abs(log(h))^(1+eps) )
# Gaussian correlation with theta=1
Rh=exp(-h^2)
  Exponential correlation with theta=1
Rhe=exp(-abs(h))
```

plot(h,bound,type='l',lwd=3,col="grey",ylim=c(0,1),xlab="h lines(h,1-Rh,col="blue") lines(h,1-Rhe,col="orange")





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- For the Gaussian correlation model, $R(h) = exp(-\theta h^2)$, then $Z(\cdot)$ is a.s. continuous and a.s. infinitely differentiable.
- For the cubic correlation model, Z(·) is a.s. continuous and 2 times differentiable.

•
$$E\left[\frac{\partial}{\partial x_i}Z(x)\right] = \frac{\partial}{\partial x_i}E\left[Z(x)\right]$$

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• $Cov\left(\frac{\partial}{\partial x_i}Z(x), \frac{\partial}{\partial x_j}Z(x')\right) = \frac{\partial^2}{\partial x_i\partial x_j}Cov\left(Z(x), Z(x')\right) = \frac{\partial^2}{\partial x_i\partial x_j}c(x-x').$

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- $\operatorname{Cov}\left(\frac{\partial}{\partial x_i}Z(x), \frac{\partial}{\partial x_j}Z(x')\right) = \frac{\partial^2}{\partial x_i\partial x_j}\operatorname{Cov}\left(Z(x), Z(x')\right) = \frac{\partial^2}{\partial x_i\partial x_j}c(x-x').$
- Clearly we at least need c(·) to be twice differentiable to obtain the correlation/covariance function of the derivative process.

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- Let T be a translation operator:

$$T({Z(t)}: t = 1, 2, ...) = {Y(t): t = 1, 2, ...}$$

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• $T^k = T(T(\ldots))$

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- For our time-indexed process analogy here, this definition says that almost every realization of the time-series
 {*Z*(*t*), *t* = 1, 2, ...} when successively translated completely
 fills the space of possible trajectories.
- i.e. the future of this ergodic process holds within it any type fo behaviour allowable under its probability measure *P*.

Ergodic Theorem[†]: Let $\{Z(t), t = 1, 2, ...\}$ be an ergodic time-series and suppose that the measureable function $g(\cdot)$ is integrable ($\int gdP$ exists, e.g. if $g = \mathcal{I}_A$ then $\int gdP = P(A)$). Then for almost every realization $\{z(t), t = 1, 2, ...\}$,

$$\lim_{n\to\infty} \frac{1}{n} \sum_{j=0}^{n-1} g(T^j(\{z(t): t=1,2,\ldots\})) = \int g dP.$$

† Birkhoff: Proof of the Ergodic Theorem (1931).

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- In practice: can we make inference of a GP observed over a fixed domain, e.g. [0,1]? This is known as *infill asympototics*.

Identifiability

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- The log-likelihood function of a GP for the observed vector $\mathbf{y} = (y_1, \dots, y_n)$ is

$$\ell(\boldsymbol{\theta}) = -\frac{n}{2}log(2\pi) - \frac{1}{2}log|\boldsymbol{\Sigma}| - \frac{1}{2}\mathbf{y}^{\mathsf{T}}\boldsymbol{\Sigma}^{-1}\mathbf{y}.$$

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• Then the parameters θ are identifiable if $\ell(\theta_1) = \ell(\theta_2)$ iff $\theta_1 = \theta_2$.

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- † For a 1-D GP with mean zero and exponential covariance $\sigma^2 exp(-\theta|h|)$ observed over the domain [0,1], then (θ, σ^2) and $\theta', \sigma^{2'}$) are **not** distinguishable with certainty from the sample path as long as $\theta\sigma^2 = \theta'\sigma^{2'}$.

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- So, only $\theta\sigma^2$ is identifiable, not the individual parameters.
- For dimension ≥ 2, with a separable covariance, this is not the case (they are identifiable).

```
source("dace.sim.r")
n=100
design=matrix(seq(0,1,length=n),ncol=1)
11=list(m1=abs(outer(design[,1],design[,1],"-")))
l.dez=list(l1=l1)
rho=0.2
alpha=1
s2s=10
rs=1/s2s
seed=sample(1:1e5,1)
set.seed(seed)
s2=1
rho=0.2 # -log(.2)*s2=1.6ish
se2=0
z0=sim.field(l.dez,rho,s2,se2=se2,alpha=alpha)
z0=z0-mean(z0)
```





set.seed(seed)
s2=1*s2s
rho=0.2
z1=sim.field(l.dez,rho,s2,se2=se2,alpha=alpha)
z1=z1-mean(z1)





set.seed(seed)
s2=1
rho=0.2^rs
z2=sim.field(1.dez,rho,s2,se2=se2,alpha=alpha)
z2=z2-mean(z2)





set.seed(seed) s2=1*s2s rho=0.2^rs # -log(rho)*s2=1.6ish z3=sim.field(1.dez,rho,s2,se2=se2,alpha=alpha) z3=z3-mean(z3)

